

CO-INTEGRATION WITH BREAKS

H_0 : NO C-I

H_1 : C-I

1. CDDW (1987)
 2. EG
 3. Johansson
 4. P.S.S (1999, 2001)
 5. Narayan (2005)
- ← Residual

All these tests no breaks or regime shifts in co-integration vectors

The second type of tests in where

a) breaks can exist

and/or

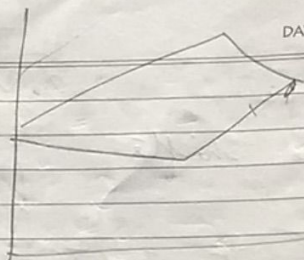
x_1, y, x_2, x_4

b) Regime shifts

TYPES OF TESTS



- 1) GREGORY AND HANSEN (1996 a, 1996 b)
- 2) JOHANSEN ETAL (2000)
- 3) Carrion-I-Sylvestre and Sanso (2006)
- 4) Luptaphol et al (2004)
- 5) Salkkahan and Luptaphol (2009)
- 6) Inoue (1999)



DATE

Existence of break does not mean regime shift

~~not correct~~ The conclusion will change of the hypothesis will change

GREGORY AND HANSEN (1996a, b)

developed Residual based test

for

H_0 : ND C.I.

H_1 : C.I.

with

a) ~~level shift~~

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b) level shift

NO break or regime shift
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c) level with trend

d) Regime switching [shift at level and slope]

There are four alternatives here

$$y_t = \alpha_1 + \alpha_2 x_t + e_t$$

$$e_t = y_t - \alpha_1 - \alpha_2 x_t$$

$$e_t = \gamma e_{t-1} + \sum_{i=1}^p \Delta B_i e_{t-i} + w$$

This marks for the first alternative: This is just repetition of EG.

Therefore, this model (alternative) becomes the null hypothesis

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$$H_0: Y_t = \mu_1 + \phi^T X_t + e_t$$

X_t = vector of exogenous variable with $I(1)$ process

$$X_t \sim I(1) \quad e_t \sim I(0)$$

This null to be tested against three alternatives.

H_{1a} : C-I with unknown single break date at level.

: with level shift at unknown date (endogenous break date)

(multivariate version of Zivot Andrews)
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if this is known then it is multivariate version of
The dummy has to be defined

$$\gamma_{t\tau} = \begin{cases} 0 & \text{if } t \leq [n, \tau] \\ 1 & \text{if } t > [n, \tau] \end{cases}$$

(gamma) $\uparrow$ unknown, tau (endogenous)

$t \in T$ scaling parameter

This is the dummy variable.

$$Y_t = \mu_1 + \mu_2 \gamma_{t\tau} + \phi^T X_t + e_t$$

μ_1 = is the level of the relation before the shift

μ_2 = is the level of the relation after the shift.

MODEL 3 [level shift with trend]

MODEL [C/T] just add the trend.

ALTERNATIVE
H_{1c}:

$$y_t = \mu_1 + \mu_2 \gamma_t + \boxed{\beta t} + \boxed{\Phi^T X_t} + e_t$$

Matrix vector

↑ trend ↑

IT MEANS

H_{1c}: C-I WITH IN PRESENCE OF LEVEL SHIFT AND TREND

MODEL 4 [REGIME SHIF]

Alternative

$$H_{1d} : y_t = \mu_1 + \mu_2 \gamma_t + \Phi_1^T X_t + \Phi_2^T X_t + e_t$$

Φ_1 = parameters before regime shift (R.S)

Φ_2 = parameters after R.S

group: Both the slope and level will have Regime

μ_1 = level parameter before R.S.

μ_2 = || || after R.S.

3
The τ test statistic developed

1) ADF

2) Z_d ← just like Phillips and Perron.

3) Z_t

He realized that, in the presence of breaks will lead to false conclusion

GHT (1996)

Here, a more comprehensive model (where regime shift and trend shift)

MODEL [C/S/T]

H_0 : NO C.I. [NULL SAME]

H_{1e} : C.I. in presence of regime shift and trend shift at unknown date

such that his model is nothing but

$$y_t = \mu_1 + \mu_2 x_{1t} + \beta_1 t + \beta_2 t$$

μ_1, β_1, θ_1 parameters before shift

μ_2, β_2, θ_2 after //

$n = 360$ (sample)

10,000 - iterations

no convergence for one week.

trend also has a dummy concept

$$y_t = \mu_1 + \mu_2 x_{1t} + \beta_1 t + \beta_2 t + \phi_1^T x_t + \phi_2^T x_t + x_3 + \epsilon_t$$

Johansen S. Mosconi R. and
Nielsen B (2001)

Peron (81) one break

Johansen (2000) — multiple breaks

Johansen's

$$\Delta W_t = \Pi W_{t-1} + u + e_t + \sum_{i=1}^{p-1} D_i W_{t-i}$$

$$\Delta W_t = u + \delta t + \sum_{i=1}^{p-1} D_i X_{t-i} + d_i \begin{bmatrix} W_{t-1} \\ \delta \end{bmatrix}$$

assuming $\Pi = \begin{bmatrix} \beta & u_2 & t \end{bmatrix}$
matrix $+ e_t$

$$\Delta W_t = \Pi W_{t-1} + \Pi_1 t + u + e_t$$

$$= \Pi = \alpha \beta'$$

$$\Pi_1 = \alpha \gamma'$$

Co-integrating
vecla

$$DW_t = \alpha \beta' w_{t-1} + \alpha \gamma' t + \mu + e_t$$

$$DW_t = \alpha (\beta' w_{t-1} + \gamma' t) + \mu + e_t$$

Coefficient
of speed of
adj.

This is the
form of the
co-integrating
vector

This is $H(r)$ model allows
for co-integration when there
is linear trend in the
model.

In pressaine schu
 β' - the rank is $n-1$

Here it should be full rank

RESTRICTED VERSION

First case

IF $\gamma = 0$
speed of
adj.

$$DW_t = \alpha (\beta' w_{t-1}) + \mu + e_t$$

only
imp.
in

S.R. still there is
a trend when γ .

However, when we meet the
Trend has died

This is called $H_{lc}(r)$ model.

Although the trend is
not there in C.I. vector

Second Case

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$\gamma = 0$ $\mu = \alpha p'$ no break
in level
and trend

$$Dw_t = \alpha (\beta' w_{t-1}) + \alpha p' + e_t$$

$$Dw_t = \alpha (\beta' w_{t-1} + p') + e_t$$

break
in level

There is no trend
here

because the factorising
 α is not of γt but
of $\alpha p'$

$H_c(r)$ model

GENERAL SETTING OF MODEL

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$$Dw_t = \begin{bmatrix} \pi_1 & \pi_2 \end{bmatrix} \begin{bmatrix} w_{t-1} \\ t \end{bmatrix}$$

$$+ \sum_{i=1}^{p-1} \Gamma_i Dw_{t-p} + e_t$$

$w_t =$ Vector of macroeconomic
variable

$$w_t = (X_t, Y_t, Z_t)$$

and assume multiple break
dates

Divide X_t vector into two samples

$$t_1 = 1 \dots q$$

$$t_2 = q+1 \dots m$$

$$\pi_0 w_{t-1} + \underbrace{\pi_1}_\text{role of trend} t$$

$i = \text{no of breaks}$

$$DW_L = [$$

DEFINE THE $H_L(r)$ HYPOTHESIS

$$H_L(r) = [\pi_0 \quad \pi_1] \begin{bmatrix} w_{t-1} \\ w_t \end{bmatrix} \leq r$$

OR

$$H_L(r) : [\pi_0 \quad \pi_1] \begin{bmatrix} w_{t-1} \\ w_t \end{bmatrix}$$

$$= \alpha \begin{bmatrix} \beta' \\ \gamma_1' \\ \vdots \\ \gamma_i' \end{bmatrix}$$

break the slope of trend function

breaks

The second matrix is changed. The break happens at trend.

$$\alpha \beta' = \pi_1' N_{t+1}$$

$$\pi_1' t = \alpha \gamma_1' + \alpha \gamma_2'$$

$$\textcircled{1} - H_L(r) : [\pi_0, \pi_1] \leq r$$

EXPAND IN λ .rank of
specific
matrix

(2) $H_L(r) = \begin{bmatrix} \pi_1 & \pi_2 & \dots & \pi_i \end{bmatrix} \leq r$

↑
rank

1st sample $\pi_1 \pi_i$
2nd $\pi_2 \pi_i$

when $\lambda = 0$

then

$$[\pi] \leq r \quad \text{because}$$

Co-integration if the inequality
is positive

$$[\pi_1 \pi_2 \dots \pi_i] = 0$$

(3) $H_{LC}(r) : [\pi] \leq r$

Same as Johansen
with no break
(1986) paper.

Finally

(4) $H_C(r) : [\pi, u_1, \dots, u_i] \leq r$

The alternative is to introduce some Dummy into the variable.

$$D_{i,t} = \begin{cases} 1 & \text{for } t = T_j - 1 \\ 0 & \text{otherwise} \end{cases}$$

it should
some
sample
because can
have my sample

summation of
dummies

$$E_{j,t} = \sum_{j'=K+1} D_{j', t-1} = \begin{cases} 1 & \text{for } T_{j'-1} + K + 1 \leq t \leq T_{j'} \\ 0 & \text{otherwise} \end{cases}$$

observation in each sample

$$t = j' - 1$$

$$j' = 2, \dots, j-1$$

K

$$\begin{cases} 1 & \text{for } T_{j'-1} + K + 1 \leq t \leq T_{j'} \\ 0 & \text{otherwise} \end{cases}$$

Vector of breaks

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$$E_t = [E_{1t}, E_{2t}, \dots, E_{qt}]$$

BASED ON ABOVE, THE MODEL

$$DW_t = \alpha \begin{bmatrix} \beta' \\ \gamma' \end{bmatrix} \begin{bmatrix} DW_{t-1} \\ tE_t \end{bmatrix} + UE_z + \sum_{i=1}^{k-1} \sqrt{i} DW_{t-1}$$

multiplication dummy

L.R. Component -

$$+ UE_z + \sum_{i=1}^{k-1} \sqrt{i} DW_{t-1}$$

$$\sum_{i=1}^k \sum_{j=2}^q k \hat{c}_{ij}^D \delta_{j,t-i} + e_t$$

From this we can

construct L.R. Hypo

$$H_L(r) = [\pi, \alpha \gamma' E_t] \leq r$$

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k, q different samples

$$H_L(r) = [\pi, UE_t] \leq r$$

$$H_L(r) = [\pi, UE_t] \leq r$$

CARRION-I-SYLVESTRE AND SANJO (2006)

allows for one break ^{in co-int.} that happens in null

H_0 : C.I with one break.

In (Gregory, Hansen) it is no co-int. and H_1 : co-int.

In this they came up with C.V of

- Known

- Unknown.

H_1 : NO C.I.

This is LM-based
based while Greg-Han
classmate is Residual based test.

DGP

$$y_t = \alpha_t + \gamma_t + X_t' \beta + \epsilon_t$$

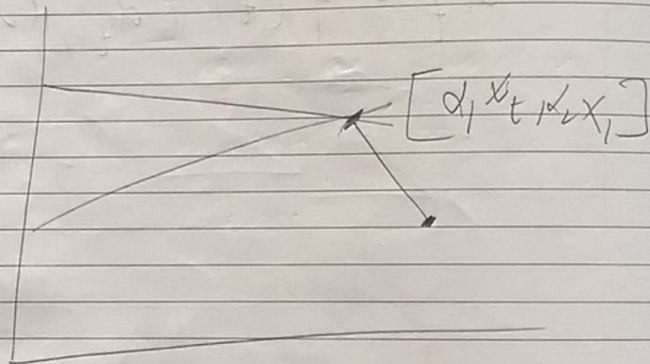
C.I. Vector

$$X_t' = X_{t-1}' + \gamma_1$$

$$\alpha_t = f(t) + \alpha_{t-1} + v_2$$

$$Dy_t = \underbrace{\mu_1}_{\text{determ}} + \underbrace{\sigma \epsilon_t \sum_{i=1}^{p-1} \Delta F_i}_{\text{Stochas}} y_{t-p} + \mu_2 + \pi y_{t-1} + \underbrace{\sigma_2 \epsilon_2}_{\text{c.v. Stochastic}}$$

↑
VAR



The co-integrating vector
is also changing

So, we have stochastic
component is also
changing.

$$x_t \sim I(1)$$

$$v_2 \sim i.i.d(0, \sigma^2)$$

$$d_t = \underbrace{f(t)}_1 + d_{t-1} + v_t$$

This is a function
indicating level
change.

Dummy

$$DT_b = \begin{cases} 1 & \text{for } t = T_b + 1 \\ 0 & \text{otherwise} \end{cases}$$

Similar

classmate ask and paper for

$$D_{tT_b} = \begin{cases} 1 & \text{for } t \geq T_b \\ 0 & \text{for otherwis.} \end{cases}$$

This is the known and
unknown here. ~~is~~ based
on ~~is~~ if you give
the break or the
data take out the
break

break
date

FIRST MODEL ~~$y_t = f(t) + \epsilon_t$~~ ODT_bCOMPLETE

$$y_t = f(t) + d_{t-1} + v_2 + \gamma_t + \beta(x_{t-1} + v_1) + \epsilon_t$$

$$= f(t) + d_{t-1} + \gamma_t + \beta' x_{t-1} + \underbrace{v_2 + \beta' v_1 + \epsilon_t}_m$$

$$y_t = f(t) + d_{t-1} + \gamma_t + \beta' x_{t-1} + m$$

$$m \sim \text{i.i.d. } N(0, \sigma^2)$$

$$x_{t-1} \sim I(1)$$

FIRST MODEL

$$Y = 0 \quad f(t) = \theta(DT_b)_t$$

$$Y_t = d_{t-1} + \theta(DT_b)_t$$

This break
is affecting at
one point

Further Restrict this
model

a) $Y \neq 0 \quad f(t) = \theta(DT_b)_t$

$$Y_t = d_{t-1} + \theta(DT_b)_t$$

not said anything
about that

$$+ \beta' \underline{x_{t-1}} + m$$

allows
level shift on
point.

Deterministic component.

$$Y_t + \beta' x_{t-1} + m$$

level shift

with trend

b) If $\gamma \neq 0$

$$f(t) = \lambda (DU_t)_t$$

$$y_t = \alpha_{t-1} + \gamma_t + \lambda (DU_t)_t + \beta' x_{t-1} + \epsilon_t$$

Model B

↓
slope of
trend func. is
changing

Next

$$c) \gamma' \neq 0 \quad f(t) = \theta (DTb)_t + \lambda (DU_t)_t$$

$$y_t = \alpha_{t-1} + \gamma_t + \theta (DTb)_t + \lambda (DU_t)_t + \beta' x_{t-1} + \epsilon_t$$

Model C: level and slope of
trend function to shift.

Hao (1996) —

Bartley ^{et al} (2001) — level shift and trend which is model A

It allows for both
deterministic and
stochastic component to
change.

Now, if we assume

$$\gamma = 0$$

$$f(t) = \theta (DT_y) t + \beta' x_t (DT_y) t$$

multiplicative Dum

trend

The model becomes

$$y_t = d_{t-1} + \theta_1 (DT_y)_t + \theta_2 X_t (DT_y)_t + \beta' X_t + e_t$$

$$= d_{t-1} + \theta_1 (DT_y)_t + \underbrace{(\theta_2 DT_y + \beta')}_{\text{matrix}} X_t + e_t$$

|
slope of the C.I. vector

This is the stochastic component

This is Model (D) $\leftarrow$ It allows for deterministic

and stochastic components
to shift (but without trend)

In the next

$$\gamma \neq 0 \quad f(t) = \theta(DT_b)_t + \beta' X_t (DT_b)_t + \lambda(Du)_t$$

Then

$$y_t = d_{t-1} + \gamma_t + \theta(DT_b)_t + \beta' X_t (DT_b)_t + \lambda(Du)_t \cdot X_t \beta' + u_t$$

$$y_t = (d_{t-1} + \theta(DT_b)_t) + (\gamma_t + \lambda(Du)_t) + \left[\left[\theta(DT_b)_t + \beta \right] X_t \right] + u_t$$

Here both the trend and deterministic component will change.

Now, set the hypothesis of ~~the~~ ^{linear} model.

$$y_t = d_t + X_t' \beta + \epsilon_t$$

$$X_t = X_{t-1} + v_t$$

$$d_t = \text{~~some~~ } f(t) + d_{t-1} + u_t$$

Restriction

$$X_k \sim I(1)$$

$$u_t$$

variance

$$H_0: \sigma_w^2 = 0 \leftarrow$$

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$$H_1: \sigma_w^2 > 0 \leftarrow (\text{ " " " " NO C.I. })$$

$\Rightarrow \sigma_w^2 = 0$ (Caveat)

then $\Delta_t = 0 = f(t) \neq \Delta_{t-1} + \frac{1}{2}$

Then the model y_t will collapse to

$$y_t = g(t) + \beta' X_t + \epsilon_t$$

FDR MODEL A

$$; g(t) = \alpha + \theta(D_{ut})t$$

where $g(t)$ is an indication function of whether break will happen

$$g(t) = \alpha + \theta (DU_t)_t$$

Null hypothesis under A_2

$$g_t^{AN} = [\alpha + \theta (DU_t)_t] + \beta' X_t + \epsilon_t$$

Intercept

$g(t)$

Dummy variable

FOR ^{DIFF.} MODEL ~~B~~

$$g(t)_A = \alpha + \gamma t + \theta (DU_t)_t + \gamma_t$$

$$g(t)_B = \alpha + \gamma t + \tau (DT_t)_t$$

$$g(t)_C = \alpha + \theta (DU_t)_t + \tau (DT_t)_t + \gamma_t$$

The form of
contention is $g(t)$

Dummy affected slope

This is
only for
deterministic
component

Dummy affect level

AFFECTING - BOTH DETERMINISTIC AND STOCHASTIC - - -

$$y_t = g(t) [X_t \beta_1' + X_t \beta_2' D u_t] e^t$$

$$g_t^D = \lambda + \theta D u_t$$

$$g_t^E = \alpha + \gamma_t + \theta (D u_t)_t + \gamma (D T_b)'_t$$